

Rank-based models with listings and delistings

Theory and calibration

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Joint work in progress with Martin Larsson, Licheng Zhang (CMU)

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Long-term modelling of financial markets



- Many market participants are interested in **long-time horizons**
 - Pension funds, Trust funds, Endowments, etc.
- We will focus on **equity markets** (stocks).
- Modelling any noisy system over a long period of time is **challenging**.
 - but **may be rewarding!**

Long-term modelling of financial markets



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- Modelling any noisy system over a long period of time is **challenging**.
 - but **may be rewarding!**

Questions:

- ① What **features** of equity markets **persist** over long-time horizons
- ② Can we **develop models** capturing such features and procedures for **statistical calibration**?

Capital distribution curve



- Let S_1, \dots, S_N denote the market capitalization processes of N companies.
- Set

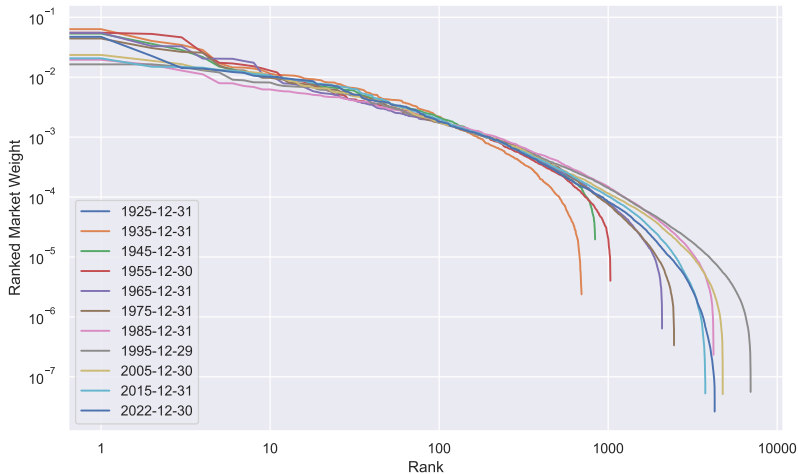
$$\mu_i(t) = \frac{S_i(t)}{S_1(t) + \dots + S_N(t)} \quad i = 1, \dots, N$$

to be the market weights.

- The ranked market weights $\mu_{(1)} \geq \mu_{(2)} \geq \dots \geq \mu_{(N)}$ are remarkably stable over time.

Capital distribution curve

Ranked market weights over time



Data Source: CRSP

First-order ranked based models



- **Curve stability** was first observed by Robert Fernholz who developed **Stochastic Portfolio Theory (SPT)** in 2002
 - SPT further developed together with Ioannis Karatzas, his students including Kardaras, Ruf and many others.

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- Fernholz proposed **rank-based** models, which are reduced-form models that can **capture the empirical stability** of the curve
 - **Advantages for calibration** due to **continuity of data**.

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- Letting $X_i = \log S_i$ the model postulates dynamics

$$dX_i(t) = \gamma_{r_i(t)} dt + \sigma_{r_i(t)} dW_i(t)$$

where $\gamma_i \in \mathbb{R}$, $\sigma_i > 0$, $r_i(t)$ is the **rank** asset i occupies at time t and W is an N -dimensional **Brownian Motion**.

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- Under the **stability condition**

$$\bar{\gamma} := \frac{1}{N} \sum_{k=1}^N \gamma_k > \frac{1}{n} \sum_{k=1}^n \gamma_k, \quad n = 1, \dots, N-1,$$

the market weights are **ergodic** representing stability of the curve.

- When $\gamma_N \gg \gamma_i$ for every other i we say the model is **Atlas-like**.

Parameter calibration



- **Ranked log-caps:** $X_{(1)} \geq X_{(2)} \geq \dots \geq X_{(N)}$.
- **Ranked dynamics:**

$$dX_{(k)}(t) = \gamma_k dt + \sigma_k d\widetilde{W}_k(t) + \frac{1}{2}d\Lambda_k(t) - \frac{1}{2}d\Lambda_{k-1}(t),$$

where Λ_k is the **local time** at zero of $X_{(k)} - X_{(k+1)}$ and $\Lambda_0 = \Lambda_N = 0$.

- This term Λ_k **activates** when two **particles collide** $X_{(k)} = X_{(k+1)}$ and ensures the ordering $X_{(k)} \geq X_{(k+1)}$ persists

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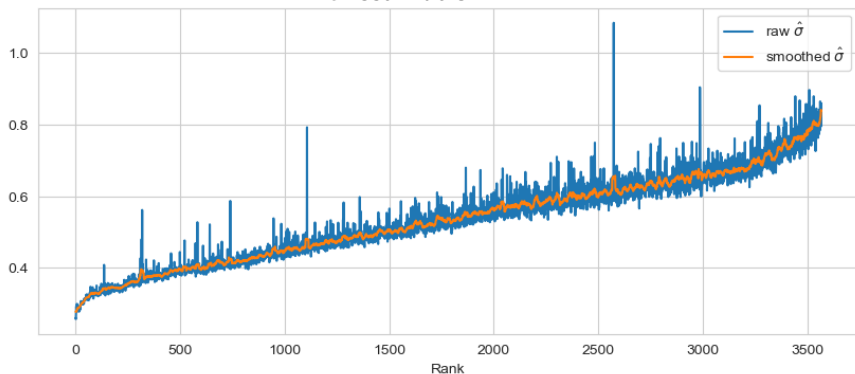
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- **Drift calibration:** $\gamma_k - \bar{\gamma} = \frac{1}{2} \lim_{T \rightarrow \infty} (\frac{1}{T} \Lambda_{k-1}(T) - \frac{1}{T} \Lambda_k(T))$.
→ Requires estimating **collision rates** $\lambda_k = \frac{1}{2} \lim_{T \rightarrow \infty} \frac{1}{T} \Lambda_k(T)$.
→ **Efficient** method using a so-called **"Master formula"** for portfolio generation available and developed in Fernholz (2002)

Calibrated Parameters

We use 51 years of CRSP data: 1973-2024 with $N = 3565$

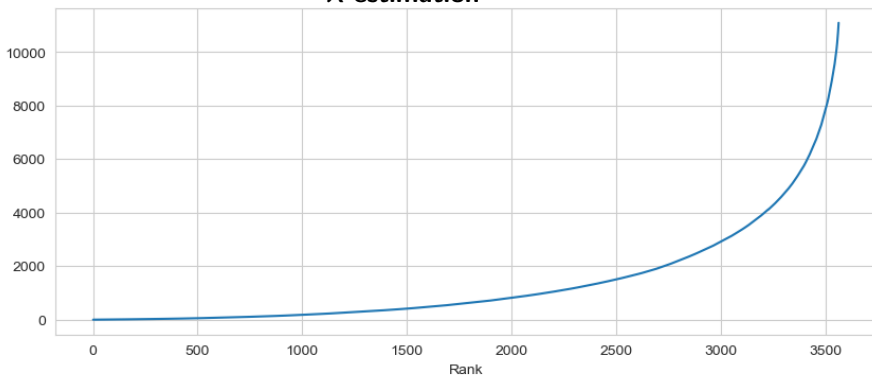
σ estimation



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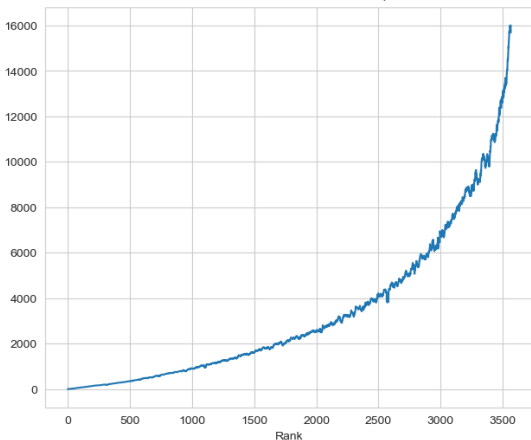
λ estimation



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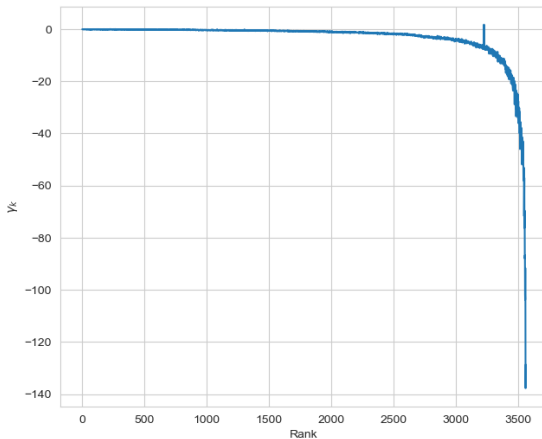
volatility normalized λ/σ^2



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γ estimation



Not plotted: $\gamma_d \approx 11000$

Puzzles



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 - Directly linked to **one-sided collision** at final rank.
 - **Artefact of the model** due to fixed number of stocks.
 - Is there a simple ad-hoc way to **modify the estimates**?
 - **NO! Removing the extreme estimate would violate stability condition of capital distribution curve.**

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Issues seem to be linked with the smallest stock/fixed universe.

- We try to address this by **allowing for listings and delistings**.
 - How prevalent are they?
- Campbell & Wong (2024) identified these as **important drivers** for capital distribution curve stability.

Model Setup



- We model a **variable equity universe** with a focus on developing **estimators for empirical calibration**,
 - Some recent literature on equity models with variable assets: Sarantsev & Karatzas (2016), Bayraktar, Kim & Tilva (2024)
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$$dX_i(t) = \gamma_{r_i(t)} dt + \sigma_{r_i(t)} dW_i(t), \quad \beta_i \leq t \leq \delta_i,$$

where $r_i(t)$ is the rank of asset i at time t **among the listed assets**.

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where $r_i(t)$ is the rank of asset i at time t **among the listed assets**.

- We set $I(t) = \{i : t \in [\beta_i, \delta_i]\}$ and $N(t) = |I(t)|$.
- We assume **nondegeneracy**: $N(t) \geq 1$ and **finite activity**:
 $\sum_{t \leq T} \Delta N(t) < \infty$ for every $T > 0$.

Ranked Dynamics

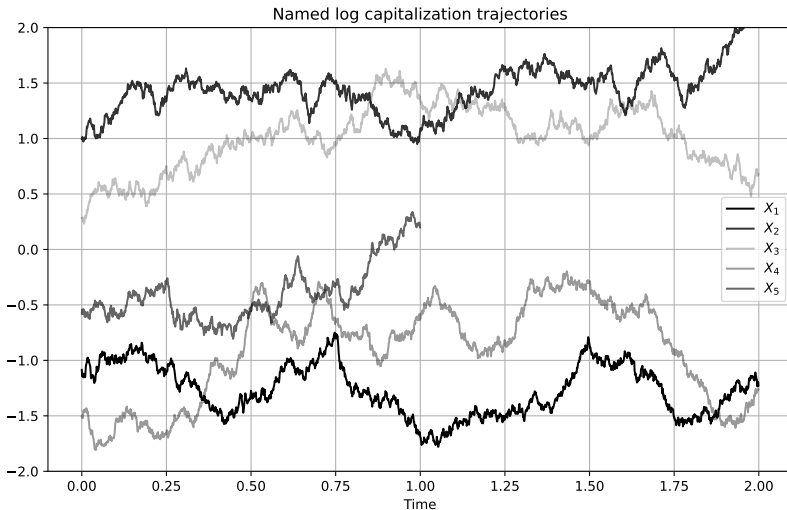


- Despite X_i being continuous on its lifetime $X_{(t)}$ experiences jumps at listing and delisting times.
- Moreover, the market capitalization $\Sigma(t) = \sum_{i \in I(t)} S_i(t)$, the market weights

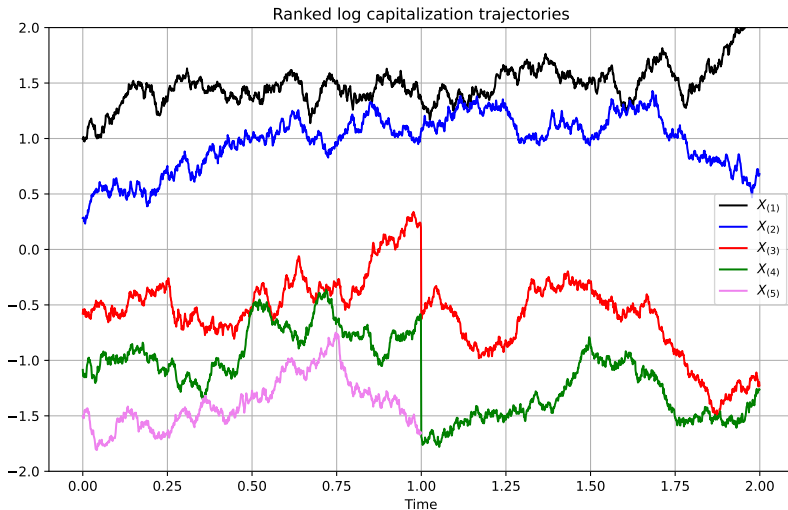
$$\mu_i(t) = \frac{S_i(t)}{\Sigma(t)} \mathbf{1}_{\{i \in I(t)\}},$$

and the ranked market weights $\mu_{(t)}$ experience jumps as well.

Illustration

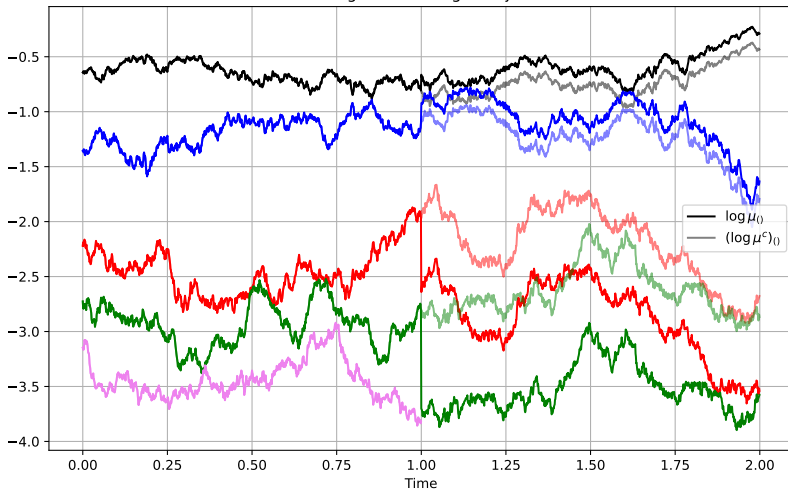


Illustration



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Ranked log market weight trajectories



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- Jumps have a larger effect on the smaller ranks.
- Care is needed dealing with portfolios and the collision estimator.
- Indeed, the wealth $V^{\mathcal{M}}$ of an investor trading the market portfolio now satisfies

$$\log V^{\mathcal{M}}(t) = \log \Sigma^c(t),$$

whereas classically $\log V^{\mathcal{M}}(t) = \log \Sigma(t)$.

Collision estimation with listings/delistings



- We developed a new “Master formula” for portfolio generation in this setting with listings and delistings.
- As in the classical setting the collision estimator can be derived by looking at the large-cap portfolio investing in the top k assets.
- In the **fixed investment universe**

$$d\Lambda_k(t) = \frac{\mu_{(1)}(t) + \dots + \mu_{(k)}(t)}{\mu_{(k)}(t)} d \log \left(\frac{V^{\mathcal{M}_k}}{V^{\mathcal{M}}} \times \frac{1}{\mu_{(1)} + \dots + \mu_{(k)}} \right) (t)$$

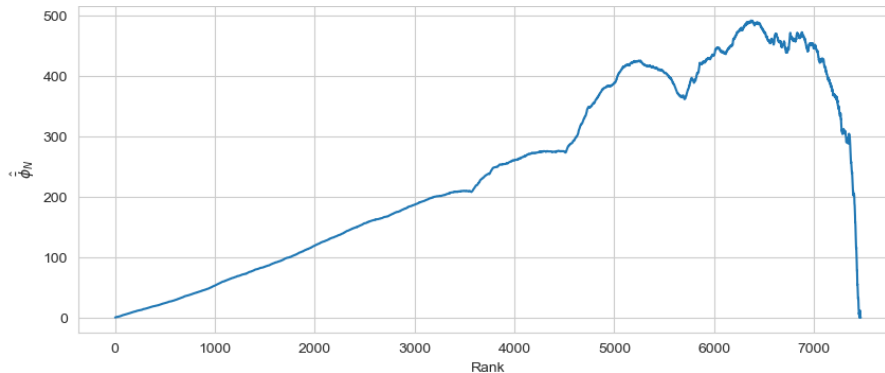
- With **listings/delistings** (on the set $\{|I(t)| \geq k\}$):

$$d\Lambda_k(t) = \frac{\tilde{\mu}_{(1)}(t) + \dots + \tilde{\mu}_{(k)}(t)}{\tilde{\mu}_{(k)}(t)} d \log \left(\frac{V^{\mathcal{M}_k}}{V^{\mathcal{M}}} \times \frac{1}{\tilde{\mu}_{(1)} + \dots + \tilde{\mu}_{(k)}} \right) (t),$$

where $\log \tilde{\mu}(t) = \log \mu^c(t)$.

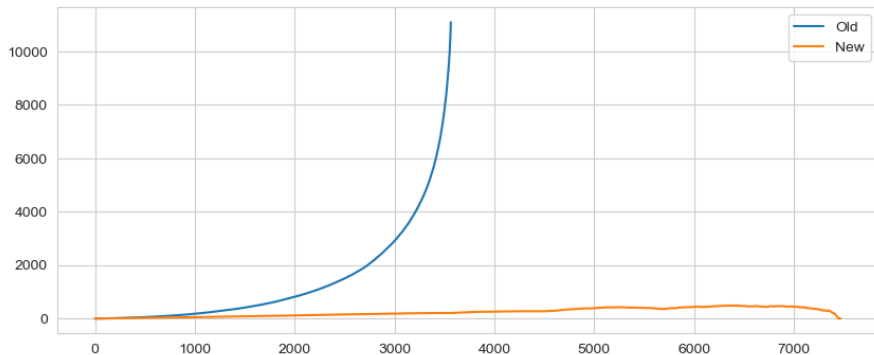
Collision estimation results

New collision estimator



Collision estimation results

A comparison



Collision estimation with listings/delistings



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where $\log \tilde{\mu}(t) = \log \mu^c(t)$.

The original estimator registers rank change caused by a listing/delisting as a collision.

- Although the **original estimator is consistent** in the fixed asset model, when applied to real data, it produces **bias**.
- The bias **propagates** effecting the **smallest stocks the most** since a listing/delisting at rank k causes a **jump for each $\mu_{(l)}$, $l > k$** .

Particle Density

- What is a **driver of collisions**?
- Naively, we expect
 - **More collisions**, if **highly volatile**,
 - **Fewer collisions** if neighbours are **positively correlated**,
 - **More collisions** if particles **tightly packed**.

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- We define average **particle density** at rank k as

$$\phi_k = \lim_{T \rightarrow \infty} \frac{1}{T} \int_0^T \frac{2n-1}{X_{(k-n)}(t) - X_{(k+n)}(t)} dt.$$

for a hyperparameter n .

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What is the relationship between σ_k, ρ_k, ϕ_k and how is this related to the collision rates λ_k ?

Local Model



- Studying this in the global model is **difficult**, so build a **local model**.
- **Idea:** For a **fixed rank k** create a **synthetic large particle model** where a typical particle behaves like our rank k one.

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- We choose the following model

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with $d[W_i, W_j](t) = \rho dt$ for all $i \neq j$.

- This is an **Ornstein–Uhlenbeck process** with **stationary measure** $N(0, \Sigma_N)$ where $\Sigma_N = (1 - \rho)I_N + \rho \mathbf{1}_{N \times N}$.

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- To study rank k we would take $\sigma = \sigma_k$, $\rho = \rho_k$, start the process at stationarity and study the **median particle** $X_{(0)}$ as the typical one.
 - The analysis leads to the **same conclusion for any fixed quantile**.

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 - The analysis leads to the **same conclusion for any fixed quantile**.
- We are interested in **understanding the relationship** between σ, ρ, λ and ϕ as $N \rightarrow \infty$.

Some order statistics results



- We can write $X_i^N = \sqrt{\rho}Y + \sqrt{1-\rho}Z_i^N$ where Y and $(Z_i^N; i = -N, \dots, N)$ are IID $N(0,1)$ random variables.
- Hence the gaps satisfy $X_i^N - X_j^N = \sqrt{1-\rho}(Z_i^N - Z_j^N)$ so it is enough to [study the IID case](#).

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Proposition (I., Larsson, Zhang (2025+))

- 1 Then density f^N on $[0, \infty)$ of the random variable $N(Z_{(0)}^N - Z_{(1)}^N)$ satisfies $\lim_{N \rightarrow \infty} f^N(0) = \sqrt{2/\pi}$.
- 2 The random variable $N(Z_{(-n)} - Z_{(n)})$ converges in distribution as $N \rightarrow \infty$ to a $\Gamma(2n, \sqrt{2/\pi})$ random variable.
- 3 The weak convergence in part two *extends* to the unbounded function $h(x) = 1/x$.

Collision parameter in the local model



- We have

$$\begin{aligned}\frac{\lambda_0^N}{N} &= \lim_{\epsilon \rightarrow 0} \mathbb{E} \left[\frac{1}{T} \int_0^T \frac{1}{N\epsilon} 1_{\{X_{(0)}^N(t) - X_{(1)}^N(t) < \epsilon\}} d[X_{(0)}^N - X_{(1)}^N](t) \right] \\ &= 2\sigma^2 \sqrt{1 - \rho} f^N(0) \rightarrow 2\sigma^2 \sqrt{1 - \rho} \times \sqrt{\frac{2}{\pi}},\end{aligned}$$

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- and

$$\begin{aligned} \frac{\phi_0^N}{N} &= \mathbb{E} \left[\frac{2n - 1}{N(X_{(-n)} - X_{(n)})} \right] \rightarrow (2n - 1) \int_0^\infty \frac{\lambda^{2n}}{\Gamma(2n)} x^{2n-2} e^{-\sqrt{2/\pi}x} dx \\ &= \sqrt{\frac{2}{\pi}} \times \frac{1}{\sqrt{1 - \rho}}. \end{aligned}$$

Collision parameter in the local model



- We have

$$\begin{aligned} \frac{\lambda_0^N}{N} &= \lim_{\epsilon \rightarrow 0} \mathbb{E} \left[\frac{1}{T} \int_0^T \frac{1}{N\epsilon} 1_{\{X_{(0)}^N(t) - X_{(1)}^N(t) < \epsilon\}} d[X_{(0)}^N - X_{(1)}^N](t) \right] \\ &= 2\sigma^2 \sqrt{1 - \rho} f^N(0) \rightarrow 2\sigma^2 \sqrt{1 - \rho} \times \sqrt{\frac{2}{\pi}}, \end{aligned}$$

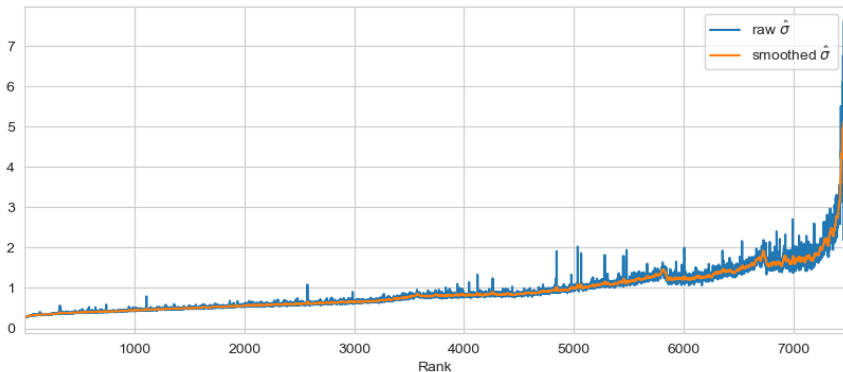
- and

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- Hence, $\phi_0^N \approx \frac{\lambda_0^N}{2\sigma^2(1-\rho)}$.
- Let see how the listing and delisting model performs with real data.

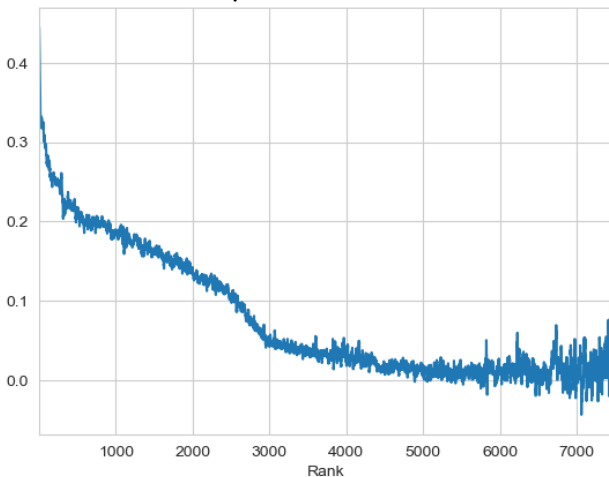
Collisions and Particle Density: Empirics

σ estimation



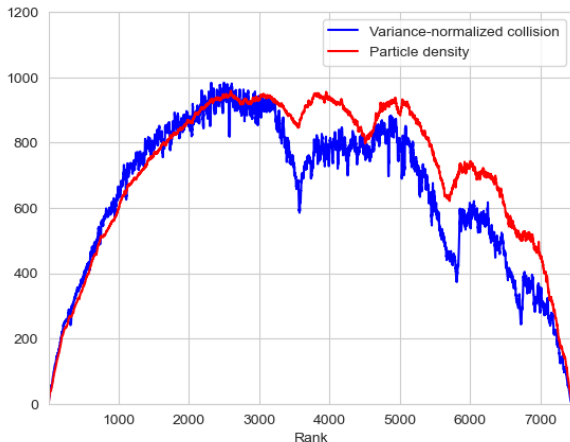
Collisions and Particle Density: Empirics

ρ -estimation



Collisions and Particle Density: Empirics

ϕ_k vs $\frac{\lambda_k}{2\sigma_k^2(1-\rho_k)}$ with $n = 5$.



A more granular look?

Conclusion and future work



Summary of results

- Introduced a **rank-based model** with **listings and delistings**,
- Derived a **new master formula** for portfolio generation
- Derived **collision estimator accounting for listings/delistings**, which **corrects bias** of previously used estimator when applied to **real data**,
- Studied **local model** and connected collisions to **particle density**.

To be done:

- Estimate **listing and delisting rates**,
- Pick a (Markovian) **birth/death mechanism** for **global model**,
- Conduct **numerical and simulation experiments** for global model.
- **Theoretical analysis** of global model?

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Thank you!

Master Formula for Functional Generation

- In this setting we say a portfolio π is **functionally generated** if

$$\log \left(\frac{V^\pi(T)}{V^{\mathcal{M}}(T)} \right) = \log \left(\frac{G(\tilde{\mu}(T))}{G(\tilde{\mu}(0))} \right) + \Gamma(T)$$

for some function $G : \bigcup_d \mathbb{R}^d \rightarrow \mathbb{R}$ and process of **finite variation** Γ .

- Here $\log \tilde{\mu}(t) = \log \mu^c(t) \rightarrow$ differs from the standard setting.

Theorem (I., Larsson, Zhang 2025+)

For a C^2 function F , the portfolio $\pi_i(t) = \sum_k \eta_k(t-) \mathbf{1}_{\{r_i(t-)=k\}}$ which invests

$$\eta_k(t) = \tilde{\mu}_{(k)}(t) \left(\partial_k \log F(\tilde{\mu}_{(l)}(t)) + \frac{1 - \sum_{\ell=1}^{N(t)} \tilde{\mu}_{(\ell)}(t) \partial_\ell \log F(\tilde{\mu}_{(l)}(t))}{\sum_{\ell=1}^{N(t)} \tilde{\mu}_{(\ell)}(t)} \right) \mathbf{1}_{\{k \leq I(t)\}}$$

in the **asset at rank k** is functionally generated by $G(x) = F(x_0)$ with

$$d\Gamma(t) = -\frac{1}{2} \sum_{k,\ell=1}^{N(t)} \frac{\partial_{k\ell} F(\tilde{\mu}(t))}{F(\tilde{\mu}(t))} d[\tilde{\mu}_{(k)}, \tilde{\mu}_{(\ell)}](t) - \frac{1}{2} \sum_{k=1}^{N(t)} (\eta_k(t) - \eta_{k+1}(t)) d\Lambda_k(t).$$

Collision rates estimation

- Applying this with the function $F(x) = x_{(1)} + \dots + x_{(k \wedge d)}$ for $x \in \mathbb{R}^d$ yields that the large-cap portfolio of size k ,

$$\begin{aligned} \pi_i(t) &= \frac{\tilde{\mu}_i(t)}{\tilde{\mu}_{(1)}(t) + \dots + \tilde{\mu}_{(k \wedge N(t))}(t)} \mathbf{1}_{\{r_i(t-) \leq k\}} \\ &= \frac{\mu_i(t-)}{\mu_{(1)}(t-) + \dots + \mu_{(k \wedge N(t-))}(t-)} \mathbf{1}_{\{r_i(t-) \leq k\}} \end{aligned}$$

has wealth process

$$\begin{aligned} \log \left(\frac{V^{\mathcal{M}_k}(T)}{V^{\mathcal{M}}(T)} \right) &= \log \left(\frac{\tilde{\mu}_{(1)}(T) + \dots + \tilde{\mu}_{(k \wedge N(T))}(T)}{\tilde{\mu}_{(1)}(0) + \dots + \tilde{\mu}_{(k \wedge N(0))}(0)} \right) \\ &\quad + \frac{\tilde{\mu}_{(k \wedge N(t))}(t)}{\tilde{\mu}_{(1)}(t) + \dots + \tilde{\mu}_{(k \wedge N(t))}(t)} d\Lambda_k(t), \end{aligned}$$

with the convention that $d\Lambda_k(t) = 0$ on $\{|I(t)| < k\}$.

Discretized Estimators

- Old estimator for local time

$$\frac{1}{T} \sum_{i=0}^{M-1} \frac{S_{(1)}(t_i) + \dots + S_{(k)}(t_i)}{S_{(k)}(t_i)} \log \left(\frac{S_{(1)}(t_i + 1) + \dots + S_{(k)}(t_i + 1)}{S_{n_1(t_i)}(t_i + 1) + \dots + S_{n_k(t_i)}(t_i + 1)} \right),$$

where $n_k(t)$ is the name occupying the k 'th rank at time t .

- New estimator for local time:

$$\frac{1}{Tp(k)} \sum_{i=0}^{M-1} \mathbf{1}_{\{|I(t_i)| \geq k\}} \frac{S_{(1)}(t_i) + \dots + S_{(k)}(t_i)}{S_{(k)}(t_i)} \times \log \left(\frac{S_{J_{(1)}^{t_i, t_i+1}}(t_i + 1) + \dots + S_{J_{(k)}^{t_i, t_i+1}}(t_i + 1)}{S_{J_{(1)}^{t_i, t_i}}(t_i + 1) + \dots + S_{J_{(k)}^{t_i, t_i}}(t_i + 1)} \right),$$

where

- $p(k) = \frac{1}{M} \sum_{i=0}^{M-1} \mathbf{1}_{\{|I(t_i)| > k\}}$,
- $J_{(\ell)}^{t,s}$ is the name of of ℓ 'th largest market cap based on time s values and out of only the names that are listed at time t and $t + 1$.